

4. Moment Generating Function Summary

- a. General discrete distribution for X with probabilities π_i on points ζ_i : $m_X(t) = \sum_i \pi_i \exp(\zeta_i t)$ for $t \in \mathfrak{R}$.
- b. Binomial $X \sim \text{Bin}(\pi, m)$: $m_X(t) = (\exp(t)\pi + 1 - \pi)^m$ for $t \in \mathfrak{R}$.
 - i. Bernoulli trial as a special case with $m = 1$.
- c. Poisson $X \sim \text{Pois}(\lambda)$: $m_X(t) = \exp([\exp(t) - 1]\lambda)$ for $t \in \mathfrak{R}$.
- d. Negative Binomial $N \sim \text{NBin}(\pi, k)$: $m_N(t) = \exp(kt)[1 - \exp(t)(1 - \pi)]^{-k} \pi^k$ if $t < -\ln(1 - \pi)$.
 - i. Geometric as a special case with $k = 1$.
- e. Gamma distribution $X \sim \Gamma(k, \beta)$: $m_X(t) = (1 - \beta t)^{-k}$ if $t < 1/\beta$.
 - i. Exponential distribution as a special case with $k = 0$.
- f. Laplace $X \sim \text{Lap}$: $m_X(t) = \frac{(1+t)^{-1}}{2} + \frac{(1-t)^{-1}}{2}$ for $|t| < 1$.
- g. Normal $Y \sim \text{N}(\mu, \sigma^2)$: $m_Y(t) = \exp(\frac{1}{2}t^2\sigma^2 + \mu t)$ for $t \in \mathfrak{R}$.
 - i. Standard normal as a special case with $\mu = 0, \sigma^2 = 1$.

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5. Higher-Dimensional Distributions

- a. Sample space consists of a region in \mathfrak{R}^k .
- i. Example: Continuous biological or economic measurement.
 - ii. Represent probabilities by integrals of a function over event sets
 - iii. Get a legitimate probability if function is non-negative and integrates to 1 over all of S .

b. distribution function: $F_{\mathbf{X}}(x_1, \dots, x_k) = \mathbf{P}(X_1 \leq x_1, \dots, X_k \leq x_k)$

c. probability function: $p_{\mathbf{X}}(x_1, \dots, x_k) = \mathbf{P}(X_1 = x_1, \dots, X_k = x_k)$

d. probability density function: $f_{\mathbf{X}}(x_1, \dots, x_k)$ such that

$$\int_A f_{\mathbf{X}}(x_1, \dots, x_k) dx_1 \cdots dx_k = \mathbf{P}((X_1, \dots, X_k) \in A)$$

e. probability density function is derivative of distribution function:

- i. If $\frac{\partial^k F_{X_1, \dots, X_k}(x_1, \dots, x_k)}{(\partial x_1 \cdots \partial x_k)}$ exists, then

$$\frac{\partial^k F_{X_1, \dots, X_k}(x_1, \dots, x_k)}{(\partial x_1 \cdots \partial x_k)} = f_{X_1, \dots, X_k}(x_1, \dots, x_k)$$

6. Marginal and Conditional Quantities

a. Marginal distribution function: If $m < k$,

$$\begin{aligned} F_{X_1, \dots, X_m}(x_1, \dots, x_m) &= \lim_{x_{m+1} \rightarrow \infty} \dots \lim_{x_k \rightarrow \infty} \mathbf{P}(X_1 \leq x_1, \dots, X_k \leq x_k) \\ &= \lim_{x_{m+1} \rightarrow \infty} \dots \lim_{x_k \rightarrow \infty} F_{X_1, \dots, X_k}(x_1, \dots, x_k) \end{aligned}$$

b. Marginal probability density function: If $m < k$,

$$\begin{aligned} f_{X_1, \dots, X_m}(x_1, \dots, x_m) &= \\ &\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_{X_1, \dots, X_k}(x_1, \dots, x_k) dx_{m+1} \dots dx_k \end{aligned}$$

c. Conditional distributions: If \mathbf{X} continuous, then

$$\begin{aligned} f_{X_1, \dots, X_m | X_{m+1}, \dots, X_k}(x_1, \dots, x_m | x_{m+1}, \dots, x_k) &= \\ &= f_{X_1, \dots, X_k}(x_1, \dots, x_k) / f_{X_{m+1}, \dots, X_k}(x_{m+1}, \dots, x_k) \end{aligned}$$

WMS: 6.3b, 6.4b, 6.6

D. Transformations of multiple random variables

1. The setup: $U = u(X, Y)$, $V = v(X, Y)$.

a. Inverses $X = x(U, V)$, $Y = y(U, V)$;

i. Unless stated otherwise, assume 1-1.

b. The objective:

i. joint distribution of U , V

ii. marginal distribution of U .

2. Transformations in the Discrete case:

a. $f_{U,V}(u, v) = f_{X,Y}(x(u, v), y(u, v))$

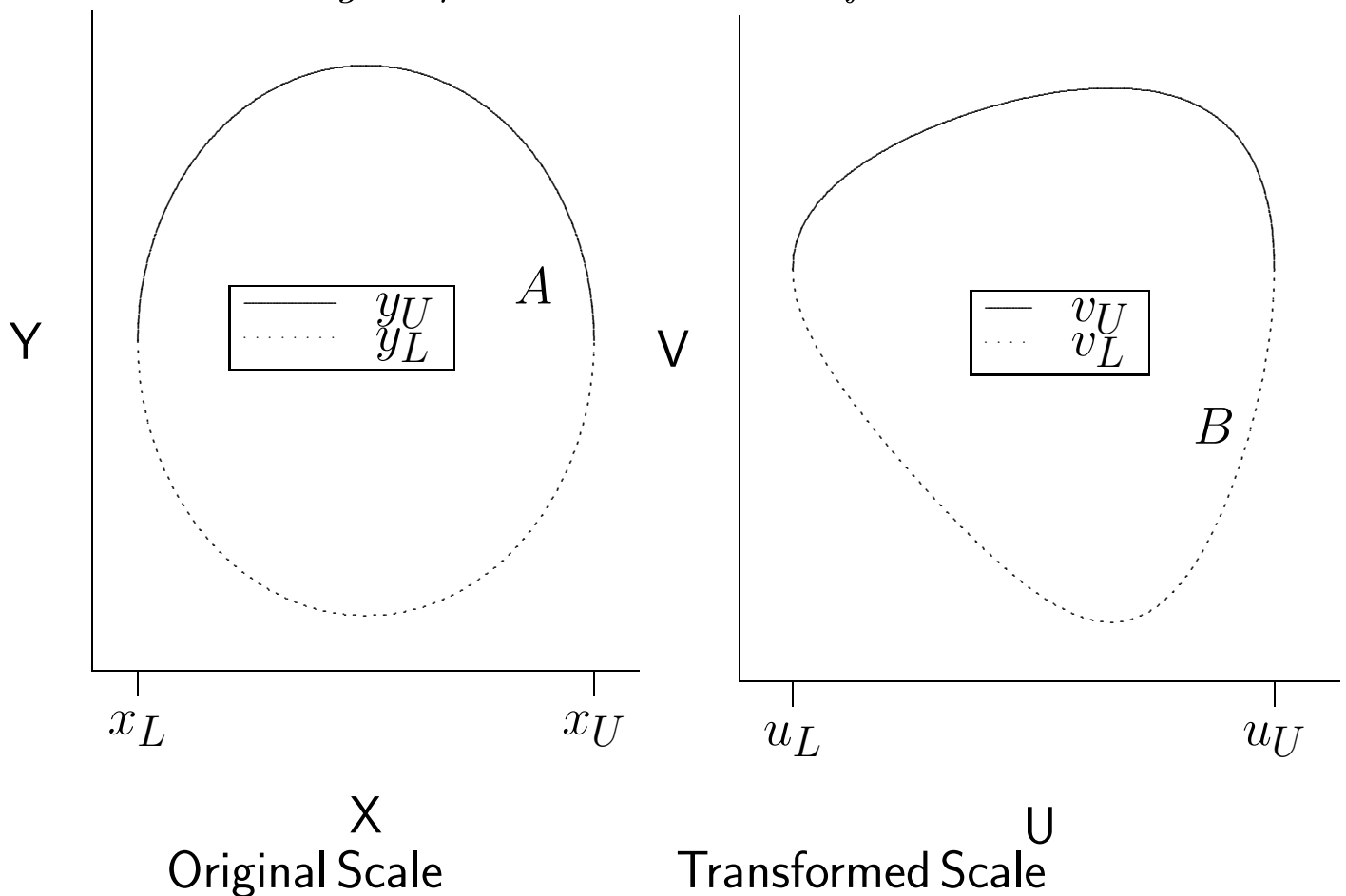
3. Transformations in the Continuous case:

a. The setup:

i. Let A be the event on X, Y scale.

ii. Let B be the event on U, V scale. See Fig. 34.

Fig. 34: Bivariate Transformation



b. $P((U, V) \in B) = P((X, Y) \in A) =$

$$\int_{x_L}^{x_U} \int_{y_L(x)}^{y_U(x)} f_{X,Y}(x, y) dy dx$$

- c. Suppose u, v differentiable with continuous partial derivatives.
- d. Let J be the Jacobian of the forward transformation.
- i. $J = \left| \det \begin{pmatrix} \frac{\partial u}{\partial x} & \frac{\partial u}{\partial y} \\ \frac{\partial v}{\partial x} & \frac{\partial v}{\partial y} \end{pmatrix} \right|$
- e. Let J^{-} be the Jacobian of the inverse transformation.
- i. $J^{-} = \left| \det \begin{pmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{pmatrix} \right| = 1/J(x(u, v), y(u, v))$
- f. Change variables to U, V : $P((\mathbf{U}, \mathbf{V}) \in B) = \int_{u_L}^{u_U} \int_{v_L(x)}^{v_U(x)} f_{X,Y}(x(u, v), y(u, v)) J^{-} du dv$
- g. Hence $f_{U,V}(u, v) = f_{X,Y}(x(u, v), y(u, v)) |J^{-}|$
- h. Multiple inverses: Sum over them.
4. Distribution of a lower-dimensional transformation
- a. Ex., (X, Y) have a joint distribution.
- b. Want distribution of $U = u(X, Y)$
- c. Change of variables formula in multiple dimensions doesn't work,
- i. because Jacobian isn't defined,
- ii. and so density formula doesn't work.
- d. Fix via transforming to a space of the same dimension and marginalizing.

- i. Choosing a convenient additional variable V
- ii. A function of X and Y
- iii. Remove its effect on the density by marginalizing.

WMS: 6.5

5. The mgf method: $U = u(X, Y)$

- a. $m_U(v) = \mathbf{E}(\exp(vu(X, Y)))$
- b. Try to recognize this.

WMS: 5.5

E. Expectations in Multiple Dimensions

1. Setup:

- a. Consider jointly-distributed random variables (X, Y)
- b. Consider some summary function $g_1(x, y)$.
- c. Want $\mathbf{E}(g_1(X, Y))$.

2. Discrete distributions:

- a. Let $Z = g_1(X, Y)$.
- b. Calculate probability function $p_Z(z) = \sum_{x \in \mathcal{X}, y \in \mathcal{Y} | z = g_1(x, y)} p_{X, Y}(x, y)$
- c. Determine sample space \mathcal{Z} .
- d. Sum $\sum_{z \in \mathcal{Z}} zp_Z(z)$

e. Equivalent to

$$\sum_{x \in \mathcal{X}, y \in \mathcal{Y}} g_1(x, y) p_{X, Y}(x, y).$$

3. Continuous case:

a. Calculate probability density function $f_Z(z)$

i. Make a second new random variable $W = g_2(X, Y)$ so that $J \neq 0$.

ii. Construct sample space \mathcal{Z} , and, for each $z \in \mathcal{Z}$, construct the sample space for W \mathcal{W}_z .

iii. Calculate inverse functions h_1, h_2 to g_1, g_2

iv. Joint probability density function is $f_{Z, W}(z, w) = f_{X, Y}(h_1(z, w), h_2(z, w)) J^{-1}$

v. Marginal probability density function is $\int f_{Z, W}(z, w) dw = \int f_{X, Y}(h_1(z, w), h_2(z, w)) J^{-1} dw$

b. Expectation is $\int_{\mathcal{Z}} z f_Z(z) dz$

i. Substitute in: $\int_{\mathcal{Z}} z \int_{\mathcal{W}_z} f_{Z, W}(z, w) dw dz = \int_{\mathcal{Z}} \int_{\mathcal{W}_z} z f_{X, Y}(h_1(z, w), h_2(z, w)) J^{-1} dw dz$

ii. Change variables back

c. Result is

$$E(Z) = \int_{\mathcal{X}} \int_{\mathcal{Y}} g_1(x, y) f_{X,Y}(x, y) dy dx.$$

i. Did not depend on our choice of g_2 .

WMS: 5.6

4. Summability:

a. Suppose X and Y have a joint probability function

$$f_{X,Y}(x, y).$$

b. Then $E(X + Y) = E(X) + E(Y)$

i. Because

$$\begin{aligned} E(X + Y) &= \iint (x + y) f_{X,Y}(x, y) dx dy \\ &= \iint x f_{X,Y}(x, y) dx dy + \iint y f_{X,Y}(x, y) dx dy \\ &= \iint x f_{X,Y}(x, y) dy dx + \iint y f_{X,Y}(x, y) dx dy \\ &= \int x f_X(x) dx + \int y f_Y(y) dy \\ &= E(X) + E(Y) \end{aligned}$$

ii. Holds for discrete variables by substituting summation for integral.

c. By extension this holds for any number of summands.

d. Ex.: a $\text{Bin}(m, \pi)$ variable has expectation $m\pi$.

e. Hence the expectation has the advantage of transforming easily.

5. Multiplicative property under independence:

a. If X and Y are independent,

b. g and h are functions

c. Then $E(g(X)h(Y)) = E(g(X)) E(h(Y))$.

i. Because

$$\begin{aligned} E(g(X)h(Y)) &= \int \int g(x)h(y)f_{X,Y}(x,y) dx dy \\ &= \int \int g(x)h(y)f_X(x)f_Y(y) dx dy \\ &= \int g(x)f_X(x) dx \int h(y)f_Y(y) dy \\ &= E(g(X)) E(h(Y)) \end{aligned}$$

ii. Holds for discrete variables by substituting summation for integral.