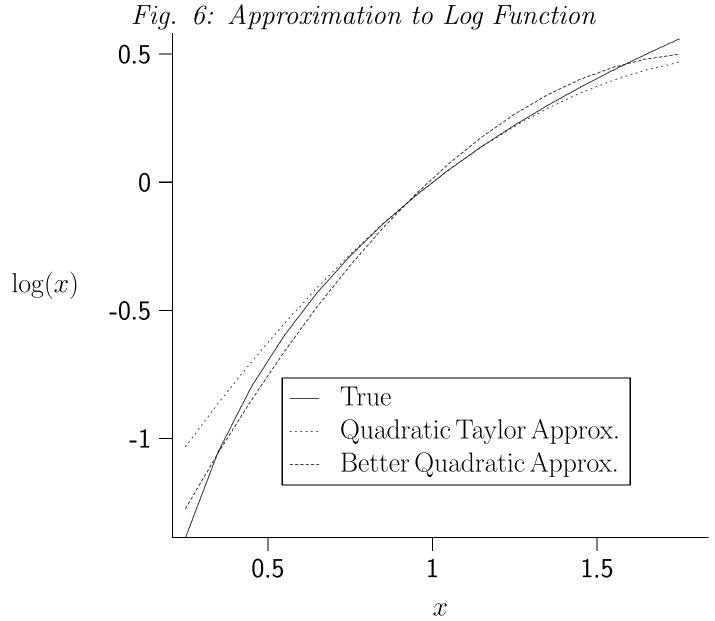
MPV: 7.1-7.2

- 9. Some explanatory variables can be transformations of existing variables.
 - a. \log , \exp , \sin , \cos can exist in a model along side the original.
 - b. More immediately, x^2 , x^3 , etc.
 - i. Result of having 1 (that is, the intercept), x, x^2, x^3, \ldots, x^k is called a polynomial of order k
 - ii. and so the model with no higher-order terms is a first order polynomial.
 - iii. Useful because well-behaved functions of the explanatory variable can be expressed as a Taylor approximation about the mean.
 - iv. Stone-Weierstrass theorem says that any continuous function on a bounded range can be approximated arbitrarily well by a polynomial.
 - v. Useful polynomials are of relatively small order.
 - c. Ex., $\log(X)$ near μ is approximately $\log(\mu) + \frac{x-\mu}{\mu} \frac{(x-\mu)^2}{2\mu^2} + O\left((x-\mu)^3\right)$.
 - i. See Fig. 6.



10. Disadvantages:

- a. Quadratic approximation uses 2 parameters to represent something that might be represented with 1 parameter times a transformation
 - i. This gets worse if you add more powers of the variable.

- b. Adding parameters allows overfitting
 - i. For certain x configurations, one can fit any n Y values exactly using $1, x, x^2, \ldots, x^{n-1}$,

ii. Design matrix
$$m{X} = \begin{pmatrix} 1 & x_1 & x_1^2 & \cdots & x_1^{n-1} \\ 1 & x_2 & x_2^2 & \cdots & x_2^{n-1} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & x_n & x_n^2 & \cdots & x_n^{n-1} \end{pmatrix}$$

- iii. X is often (but not always) non-singular.
 - Will be singular if x_i are repeated.

iv.
$$(X^{\top}X)^{-1} = X^{-1}X^{-1\top}$$

v.
$$\hat{\boldsymbol{Y}} = \boldsymbol{X} (\boldsymbol{X}^{\top} \boldsymbol{X})^{-1} \boldsymbol{X}^{\top} \boldsymbol{Y} = \boldsymbol{Y}$$
 .

- vi. This is to some extent a strawman argument, since no practical statistician adds this many terms.
- c. $X^{\top}X$ has number of rows and columns equal to the number of parameters in the model.
 - i. Large matrices, even if invertible, may be close to singular
 - ii. Loosely,
 - distance from singularity is referred to as the matrix's condition, and
 - matrices close to singular are called ill conditioned.

- iii. III-conditioned matries are bad:
 - Exact inverse leads to highly-variable responses.
 - Numerical inverse harder to compute exactly.
 - Centering variable before raising to power can help this.
- d. Because these higher-order terms are highly variable, extrapolation in this case is a bigger problem than in the linear case.
- 11. Fits are invariant to affine transformations of regressors used in polynomials
 - a. Suppose $\hat{Y}_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2$
 - b. $x_i = \gamma_1 z_i + \gamma_0$
 - c. Then

$$\hat{Y}_{i} = \beta_{0} + \beta_{1}(\gamma_{1}z_{i} + \gamma_{0}) + \beta_{2}(\gamma_{1}^{2}z_{i}^{2} + 2\gamma_{1}\gamma_{0}z_{i} + \gamma_{0}^{2})$$

$$= (\beta_{0} + \beta_{1}\gamma_{0} + \beta_{2}\gamma_{0}^{2}) + (\beta_{1}\gamma_{1} + 2\beta_{2}\gamma_{0}\gamma_{1})z_{i} + \beta_{2}\gamma_{2}^{2}z_{i}^{2}$$

$$= \alpha_{0} + \alpha_{1}z_{i} + \alpha_{2}z_{i}^{2}$$

- i. For $\alpha_0=\beta_0+\beta_1\gamma_0+\beta_2\gamma_0^2$, $\alpha_1=\beta_1\gamma_1+2\beta_2\gamma_0\gamma_1$, $\alpha_2=\beta_2\gamma_2^2$.
- d. Hence model using quadratic in $\,x_i\,$ and model using quadratic in $\,z_i\,$
 - i. give same fits.

- ii. can convert back and forth without refitting.
- e. Works only if you don't skip powers.
- f. Text calls such models hierarchical
- g. If the range of the transformed variable is small relative to the curvature of the transformation,
 - i. the higher-order terms may be almost colinear with the linear terms.
 - ii. Can mask significance of lower-order terms.
- 12. Changes interpretation of parameter estimates
 - a. Columns of design matrix cannot be treated as changable independently.
 - b. Hence in the case of polynomial terms, coefficients no longer represent the change in response associated with a unit response in the explanatory variable.
 - c. In model $\mathrm{E}\left[Y_j\right]=\beta_0+\beta_1x_j+\beta_2x_j^2$, $d\mathrm{E}\left[Y\right]/dx=\beta_1+2\beta_2x$, and is hence dependent on x .

MPV: 7.2.3

13. Trigonometric terms

a. $\delta_1 \sin(x) + \gamma_1 \cos(x)$.

- i. x should be scaled to make period 2π
- ii. Angles are measured in radians.
- b. Counterpart of higher-order terms for polynomials:

$$\delta_j \sin(jx) + \gamma_j \cos(jx)$$

- There is a counterpart to the Stone-Weierstrass theorem demonstrating that one can approximate a bounded function arbitrarily closely with trig terms.
- ii. Typically one uses only a few such terms.
- c. Careful: if you have a time scale suggesting periodicity, you probably have dependence between temporally similar observations.
- d. Terms can represent phase shift using Sum of Angles formula.

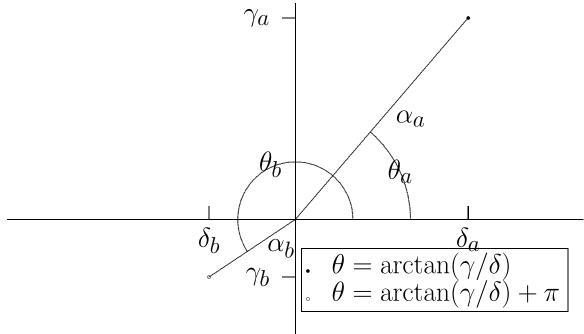
i. Let
$$\alpha=\sqrt{\delta_1^2+\gamma_1^2}$$
, $\theta=\tan^{-1}(\gamma_1/\delta_1)$, ($\theta\in(\pi/2,3\pi/2)$ if $\delta_1<0$).

- ii. Then $\delta_1 = \alpha \cos(\theta)$, $\gamma_1 = \alpha \cos(\theta)$.
 - See Fig. 7.
- iii. Then $\delta_1 \sin(x) + \gamma_1 \cos(x) = \alpha \sin(x + \theta)$.
- iv. θ is time shift.

MPV: 7.2.2

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Fig. 7: Geometry behind Trigonometric Transformation



14. Spline:

- a. A way to draw a smooth curve between two points x_0 and x_N :
 - i. Pick N-1 intermediate points $x_1 < x_2 < \cdots < x_{N-2} < x_{N-1}$ (called knots).
 - ii. Define a polynomial of degree M between x_{j-1} and x_j
- iii. Constrain so that the derivatives of order up to $\,M-1\,$ match up at knots.
- b. Use to fit pairs of points $(X_1,Y_1),\ldots,(X_n,Y_n)$.
 - i. Taken to an extreme, if all $\,X_j\,$ are unique, then we can fit all $\,n\,$ points with a polynomial of degree $\,n-1$.
- c. Denote fit by $\hat{\mu}(x)$

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- d. Choose to minimize $\sum_{j=1}^{n} (Y_j \hat{\mu}(X_j))^2$
 - i. Or penalize, to minimize $\sum_{j=1}^n (Y_j \hat{\mu}(X_j))^2 + \lambda \int_{X_{(1)}}^{X_{(n)}} \hat{\mu}''(x) \, dx$
- e. Alternative: The B-spline gives rescaled versions of the piecewise functions.
 - i. Divide by local product of knot spacing.

MPV: 7.4

- 15. Can insert polynomials with multiple explanatory variables.
 - a. Earlier ideas about hierarchical models hold here too.
 - b. If you want a model that gives the same fit under affine transformation of all regressors, you can include interaction terms
 - c. That is, the model including terms $\,x_1^2\,$ and $\,x_2^2\,$ might include $\,x_1x_2\,$ as well.
 - d. This logic is less compelling that for one variable.

MPV: 7.5

- 16. One may use orthogonal polynomials to remove colinearity
 - a. Calculate the constant, linear, quadratic, $et\ cetera.$ terms as before.

- i. Let the result be $oldsymbol{X}$
- b. Use orthogonalization as we did earlier to give orthogonal regressors
 - i. Let the result be Z
- c. Normalize if desired, to make $\sum_i z_{ij}^2 = 1$ for all i
- d. Just as before, column j of \boldsymbol{Z} is a linear combination of columns $1,\ldots,j$ of \boldsymbol{X} .
- e. Hence get same fitted values.
- f. With multiple variables, orthonormalization is applied only to the portion of the matrix corresponding to powers of one variable.

MPV: 7.3

- J. Nonparametric Regression
 - 1. Kernel smoothing:
 - a. Get an expression that is explicit rather than implicit:

$$\hat{g}(x) = \sum_{j=1}^{n} Y_j w((x - X_j)/\Delta) / \sum_{j=1}^{n} w((x - X_j)/\Delta)$$
.

- b. Weight function can be
 - i. the same as above
 - ii. Often a normal density.

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- iii. Often uniform density centered at 0.
- c. Method is kernel smoothing, and specifically is Nadaraya-Watson smoothing.
- 2. A local regression smoother has smaller bias than kernel smoother.
 - a. $\hat{g}(x) = \sum_{\ell=0}^{L} \hat{\beta}_{\ell} x^{\ell}$, for L=1

b.
$$\hat{\boldsymbol{\beta}} = \operatorname{argmin} \left(\sum_{j=1}^{n} \left(Y_j - \sum_{\ell=0}^{L} \hat{\beta}_{\ell} X_j^{\ell} \right)^2 w \left(\frac{x - X_j}{\Delta_n} \right) \right)$$
.

- 3. LOESS
 - a. f(x) fitted value at x for low-degree (viz., linear or quadratic) regression of points with X_j near x.
 - b. Specify the number of points k
 - c. Upweight points near $\,x\,$ and downweight them away from $\,x\,$
 - d. Weighting function scaled to make point in neighborhood farthest from x have weight going down to zero.
 - i. This keeps the curve smooth as x moves.
 - e. Common weight function is $w(x) = (1 |x|^3)^3$.
 - f. So $\hat{f}(x) = \hat{\beta}_0 + \hat{\beta}_1 x + \hat{\beta}_2 x^2$ for
 - i. $\hat{\beta} = \operatorname{argmin}(\sum_{j \in N(x)} (Y_j \beta_0 \beta_1 X_j \beta_2 X_j^2)^2 w((x X_j)/\Delta))$ for

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- ii. $N(x) = \operatorname{indices} \operatorname{of} \, k \, \operatorname{closest} \operatorname{points} \operatorname{to} \, x$, and
- iii. $\Delta = \max\{|X_j x| | j \in N(x)\}$.
- g. Procedure formerly Lowess, Locally Weighted Sum of Squares.
- h. Result can not be expressed as a simple formula.

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