Multivariate Tail Probability Approximations

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Kolassa and Lee

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p-value for intersection union test is calculated from $\mathsf{P}\left[\bar{X}^1 \geq \bar{x}^1 \text{ or } \bar{X}^2 \geq \bar{x}^2 | \text{other sufficient statistics} \right].$

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6 Let
$$ar{m{X}} = \sum_{i=1}^n m{X}_i / n$$



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- Approximate joint tail probabilities for \bar{X} .
- **②** Use existence of cumulant generating function to obtain relative error behavior uniform for values of \bar{x} in an open ball about the mean.
 - Measure theoretical error behavior in terms of inverse powers of \sqrt{n} .

• Density
$$f_n(\bar{w}^1, \dots, \bar{w}^p)$$
 is

$$\frac{n^p}{(2\pi i)^p} \oint \exp(n[K(\tau_1, \dots, \tau_p) - \sum_{i=1}^p \bar{w}^i \tau_i]) \ d\tau_1 \cdots d\tau_p$$

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∮ = ∫_{c1-i∞}^{c1+i∞} ··· ∫_{cp-i∞}^{cp+i∞}
c = (c1,..., cp) in the interior of the domain of K.
To find the tail probability P [\$\vec{\mathcal{X}}{\mathcal{X}} ≥ \$\vec{\mathcal{x}}{\mathcal{x}}\$] = P [\$\vec{\mathcal{X}}{\mathcal{X}} ≥ \$\vec{\mathcal{x}}{\mathcal{x}}\$],
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 Write Φ(z) = P [Z¹ ≥ z¹ ∩ ... ∩ Z^p ≥ z^p] for Z_i independent standard normals.

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- Call such an approximation *reflexive*.

Univariate Methods: Robinson (1982) approach approach Recall P $[\bar{X} \ge \bar{x}] = \frac{1}{2\pi i} \oint \exp(n[K(\tau) - \bar{x}\tau]) \frac{d\tau}{\tau}$.

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2 Re-parameterize the inversion integral in terms of ω satisfying $(\omega - \hat{\omega})^2/2 = K(\tau) - \tau x - K(\hat{\tau}) + \hat{\tau} x.$

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() Reflexive: Holds without regards to $\hat{\tau} > 0$.

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∃ →

- The tail probability $\mathsf{P}\left[\bar{\boldsymbol{X}} \geq \bar{\boldsymbol{x}}\right]$ is $\frac{1}{(2\pi i)^p} \oint \exp(n[\mathcal{K}(\tau_1, \dots, \tau_p) - \sum_{i=1}^p \bar{\boldsymbol{x}}^i \tau_i]) \frac{d\tau_1 \cdots d\tau_p}{\tau_1 \cdots \tau_p}.$
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• Integral of resulting quartic term is O(1/n).

Kolassa and Lee

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$$+\sum_{j,k,\ell,m} \mathcal{K}^{jk\ell m}(\tau^{?})(\tau_{j}-\hat{\tau}_{j})(\tau_{k}-\hat{\tau}_{k})(\tau_{\ell}-\hat{\tau}_{\ell})(\tau_{m}-\hat{\tau}_{m})/24.$$

- S Expand exp of cubic and quartic terms.
- Integral of resulting quartic term is O(1/n).
- **o** Integrate terms up to cubic term-wise.

Kolassa and Lee



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Lots of terms

Iterms are not particularly interpretable.

Lots of terms

- **2** Terms are not particularly interpretable.
 - (Multivariate) normal tail probability is multiplied by an exponential factor.

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- Iterms are not particularly interpretable.
 - (Multivariate) normal tail probability is multiplied by an exponential factor.
- You are stuck with $\hat{\tau}_j > 0$
- Ont reflexive (with definition extended beyond univariate).
 - Get around this using Boole's Law

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Wang (1990)

Shew parameterization $\sum_{j} \hat{\omega}_{j}^{2}/2 = K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j},$ $\sum_{j} (\omega_{j} - \hat{\omega}_{j})^{2}/2 = K(\tau) - \sum_{j} \tau_{j} \bar{x}^{j} - (K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j})$

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Wang (1990)

Over the set of t

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Wang (1990)

2 New parameterization $\sum_{j} \hat{\omega}_{j}^{2}/2 = K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j}$, $\sum_{j} (\omega_{j} - \hat{\omega}_{j})^{2}/2 = K(\tau) - \sum_{j} \tau_{j} \bar{x}^{j} - (K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j})$

- ω_j depends only on τ_1, \ldots, τ_j
- **2** $\hat{\omega}_j$ depends only on $\bar{x}^j, \ldots,$

Wang (1990)

Integral is

$$\frac{1}{(2\pi i)^p}\oint \exp(n\sum_j \omega_j^2 - \sum_{i=1}^p \hat{\omega}_j \omega_j)\lambda \ \frac{d\omega_1 \cdots d\omega_p}{\omega_1 \cdots \omega_p}.$$

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Wang (1990)

Integral is

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 $\bullet \ \lambda = \frac{\omega_1 \cdots \omega_p}{\tau_1 \cdots \tau_p} \frac{d\tau_1 \cdots d\tau_p}{d\omega_1 \cdots d\omega_p}.$

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2 New parameterization $\sum_{j} \hat{\omega}_{j}^{2}/2 = K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j}$, $\sum_{j} (\omega_{j} - \hat{\omega}_{j})^{2}/2 = K(\tau) - \sum_{j} \tau_{j} \bar{x}^{j} - (K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j})$ **a** ω_{j} depends only on $\tau_{1}, \ldots, \tau_{j}$ **a** $\hat{\omega}_{j}$ depends only on $\bar{x}^{j}, \ldots,$

Integral is

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() In one dimension, zeros in denominator of λ correspond to zeros in the numerator.
Lugannani and Rice Analog:

Wang (1990)

Shew parameterization $\sum_{j} \hat{\omega}_{j}^{2}/2 = K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j}$, $\sum_{j} (\omega_{j} - \hat{\omega}_{j})^{2}/2 = K(\tau) - \sum_{j} \tau_{j} \bar{x}^{j} - (K(\hat{\tau}) - \sum_{j} \hat{\tau}_{j} \bar{x}^{j})$ $\omega_{j} \text{ depends only on } \tau_{1}, \dots, \tau_{j}$ $\hat{\omega}_{i} \text{ depends only on } \bar{x}^{j}, \dots,$

Integral is

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- \blacksquare In one dimension, zeros in denominator of λ correspond to zeros in the numerator.
- 2 This fails for p > 1.

• Re-parameterize to $\xi_j = \omega_j - \sum_{k < j} \omega_k \zeta_j^k(\omega_1, \dots, \omega_{j-1})$ so that $\xi_j = 0$ if and only if $\tau_j = 0$.

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 - Ex. $\xi_1 = \omega_1$, $\xi_2 = \omega_2 \omega_1 \zeta_2^1(\omega_1)$ for $\zeta_1^2 = \omega_2(\tau_1, 0)/\omega_1$.

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 - **0** Ex. $\xi_1 = \omega_1$, $\xi_2 = \omega_2 \omega_1 \zeta_2^1(\omega_1)$ for $\zeta_1^2 = \omega_2(\tau_1, 0)/\omega_1$.
 - **2** Adjustment also makes $\{\bar{X}^1 \ge \bar{x}^1 \cap \dots, \bar{X}^p \ge \bar{x}^p\} \approx \{\hat{\Xi}^1 \ge \hat{\xi}^1 \cap \dots, \hat{\Xi}^p \ge \hat{\xi}^p\}.$

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 - **1** Ex. $\xi_1 = \omega_1$, $\xi_2 = \omega_2 \omega_1 \zeta_2^1(\omega_1)$ for $\zeta_1^2 = \omega_2(\tau_1, 0)/\omega_1$.
 - 2 Adjustment also makes $\{\bar{X}^1 \ge \bar{x}^1 \cap \dots, \bar{X}^p \ge \bar{x}^p\} \approx \{\hat{\Xi}^1 \ge \hat{\xi}^1 \cap \dots, \hat{\Xi}^p \ge \hat{\xi}^p\}.$
 - **3** Adjustment makes singularities in λ removable.







• P $[X^1 \ge x^1 \cap \ldots \cap X^p \ge x^p] \approx \bar{\Phi}(\sqrt{n}\hat{\xi}; \Sigma) + \sum_{j=1}^p \bar{\Phi}(\sqrt{n}\hat{\xi}_{-j}; \Sigma_j)\phi(\sqrt{n}\xi_j)(1/\xi_j - 1/(\hat{\tau}_j/\sigma_j)) + C.$

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• P $[X^1 \ge x^1 \cap \ldots \cap X^p \ge x^p] \approx \overline{\Phi}(\sqrt{n}\hat{\xi}; \Sigma) + \sum_{j=1}^p \overline{\Phi}(\sqrt{n}\hat{\xi}_{-j}; \Sigma_j)\phi(\sqrt{n}\xi_j)(1/\xi_j - 1/(\hat{\tau}_j/\sigma_j)) + C.$ • Σ is generated by transformation from ω to ξ .

P [X¹ ≥ x¹ ∩ ... ∩ X^p ≥ x^p] ≈ Φ(√nξ; Σ) + Σ^p_{j=1} Φ(√nξ_{-j}; Σ_j)φ(√nξ_j)(1/ξ_j - 1/(τ̂_j/σ_j)) + C. Σ is generated by transformation from ω to ξ. σ_j is a standard error for τ̂_j.

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P [X¹ ≥ x¹ ∩ ... ∩ X^p ≥ x^p] ≈ Φ(√nξ̂; Σ) + Σ^p_{j=1} Φ(√nξ̂_{-j}; Σ_j)φ(√nξ_j)(1/ξ_j - 1/(τ̂_j/σ_j)) + C. Σ is generated by transformation from ω to ξ. σ_j is a standard error for τ̂_j.

2 In order to make λ have removable singularities,

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• exponent in inversion integral is no longer exactly quadratic.

P [X¹ ≥ x¹ ∩ ... ∩ X^p ≥ x^p] ≈ Φ(√nξ̂; Σ) + ∑_{j=1}^p Φ(√nξ̂_{-j}; Σ_j)φ(√nξ_j)(1/ξ_j - 1/(τ̂_j/σ_j)) + C. Σ is generated by transformation from ω to ξ. σ_j is a standard error for τ̂_j. In order to make λ have removable singularities, exponent in inversion integral is no longer exactly quadratic.

O adjusts for this.

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• Avoids requirement that $\hat{\tau}_j > 0$.

Example

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Example

p = 2, X₁ = Z₁ + Z₂, X₂ = Z₁ + Z₃, Z_j independent exponentials.

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Example

● *p* = 2, 2 $X_1 = Z_1 + Z_2$, $X_2 = Z_1 + Z_3$, Z_i independent exponentials. **3** n = 1!

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Relative Error: Normal on raw scale



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Relative Error: Normal on SRLLR scale



Irnorm

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Relative Error: SRLLR scale, adjust for marginal non-normality



rawlr

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Relative Error: SRLLR scale, all adjustments



final

• Addressing removable singularities in C.

- Addressing removable singularities in *C*.
- **②** *p* > 2.

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- Addressing removable singularities in C.
- 2 p > 2.
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